

Analysis of non-stationary Data

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- <u>Motivation</u>
 - global stationarity, local stationarity, examples
- Non-Stationarity for continuous state space processes:
 - Fuzzy Clustering with Regression Models (FCRM)
 - *Finite Element Clustering* for continuos state space processes
- Non-Stationarity for discrete state space processes:
 - Kernel Filtering Methods
 - Single Trend Model
 - *FEM-Clustering* of Markov-chain output
 - Example I: analysis of historical *weather patterns*
 - Example II: analysis of the *historical temperatures* (1947-2007)



Non-Stationarity





Meteorology/Climate



Fluid Mechanics





Computational Finance

Biophysics/Drug Design

(Local) weak stationarity: (local) time independence of mean values and covariances

Not fulfilled in many cases

Non-Stationarity in Continuous State Space Time Series



Example: weather data analysis



Global Historical Data

Local Historical Data

Deviation from the Mean Temperature



Non-stationarity because of the climate change => standard data-analysis methods are *not applicable*







Geometrical distance: $\theta_i \in \Psi$ - time-independent cluster centers

$$g(x, \theta_i) = || x - \theta_i ||^2,$$

$$t_j, j = 1, \dots, n \in [0, T]$$

$$\sum_{i=1}^{K} \sum_{j=1}^{n} \gamma_i(t_j) || x_{t_j} - \theta_i ||^2 \rightarrow \min_{\Gamma(t),\Theta} \qquad (Bezdek1981, Hoppner et.al. 1999)$$

Iteration number (l):

$$\begin{split} \gamma_i^{(l)}(t_j) &= \begin{cases} 1 & i = \arg\min \| x_{t_j} - \theta_i^{(l-1)} \|^2, \\ 0 & \text{otherwise}, \end{cases} \\ \theta_i^{(l)} &= \frac{\sum_{j=1}^n \gamma_i^{(l)}(t_j) x_{t_j}}{\sum_{j=1}^n \gamma_i^{(l)}(t_j)}. \end{split}$$

Assumption: time-independence of cluster centers **____** <u>local stationarity</u>

Non-stationary Extension of K-Means: FCRM



Geometrical distance: time-dependent cluster centers as linear

combinations of *basis functions* $\phi_{k}(t), k = 0, \dots, \mathcal{R}$

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Geometrical distance: time-dependent cluster centers as linear

combinations of *basis functions* $\phi_k(t), k = 0, \dots, \mathcal{R}$

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$$\theta_{i}(t) = \sum_{k=0}^{\mathcal{R}} \theta_{ik} \phi_{k}(t_{j})$$
 (Hathaway and Bezdek1993)
$$\sum_{i=1}^{\mathbf{K}} \sum_{j=1}^{n} \gamma_{i}^{m}(t_{j}) \parallel x_{t_{j}} - \sum_{k=0}^{\mathcal{R}} \theta_{ik} \phi_{k}(t_{j}) \parallel^{2} \rightarrow \min_{\Gamma(t),\Theta}$$

Exercise 1: Derive an expression for the optimal estimator of

the regression parameters θ (x and γ are fixed). What happens

to this estimate if R is growing? How to define the optimal R?

Suggest and discuss the possible numerical solutions.



Time



Regularized clustering functional:

$$\mathbf{L}(\Theta, \Gamma(t)) = \int_{0}^{T} \sum_{i=1}^{\mathbf{K}} \gamma_{i}(t) g\left(x_{t}, \theta_{i}\right) \to \min_{\Gamma(t), \Theta},$$
$$\mathbf{L}^{\epsilon}(\Theta, \Gamma(t), \epsilon^{2}) = \mathbf{L}(\Theta, \Gamma(t)) + \epsilon^{2} \sum_{i=1}^{\mathbf{K}} \int_{0}^{T} \left(\partial_{t} \gamma_{i}\left(t\right)\right)^{2} dt \to \min_{\Gamma(t), \Theta}$$

$$g(x_t, \theta_i) = \| x_t - \sum_{k=0}^{\infty} \theta_{ik} \odot \phi_k(t) \|^2$$





terative Subspace Minimization

sparse

QP

can be used

$$\tilde{\mathbf{L}}^{\epsilon} = \sum_{i=1}^{\mathbf{K}} \left[a^{\mathbf{T}}(\theta_i) \bar{\gamma}_i + \epsilon^2 \bar{\gamma}_i^{\mathbf{T}} \mathbf{H} \bar{\gamma}_i \right] \to \min_{\bar{\gamma}_i, \Theta}$$

$$\sum_{i=1}^{\mathbf{K}} \tilde{\gamma}_i^{(k+1)} = 1, \quad \forall k = 1, \dots, N,$$

$$\tilde{\gamma}_i^{(k+1)} \ge 0, \quad \forall k = 1, \dots, N; i = 1, \dots, \mathbf{K}.$$

$$a(\theta_i) = \left(\int_{t_1}^{t_2} v_1(t) g(x_t, \theta_i) dt, \dots, \int_{t_{N-1}}^{t_N} v_N(t) g(x_t, \theta_i) dt \right)$$

 $g(x_t, \theta_i) = \| x_t - \sum_{k=0}^{\mathcal{R}} \theta_{ik} \odot \phi_k(t) \|^2$



Toy Example:

$$\begin{aligned} x_j(t) &= \theta_{i(t)j}(t - \bar{t}_j) + \sigma \mathbf{N}(0, 1), \quad i = 1, 2, \quad j = 1, 2, 3 \\ \theta_1 &= (0.01 \ -0.01 \ 0.01), \quad \theta_2 = (-0.01 \ 0.01 \ -0.01) \\ \bar{t} &= (0 \ 300 \ 600) \end{aligned}$$





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$$\mathbf{L}^{\epsilon}(\Theta, \Gamma(t), \epsilon^{2}) = \mathbf{L}(\Theta, \Gamma(t)) + \epsilon^{2} \sum_{i=1}^{\mathbf{K}} \int_{0}^{T} \left(\partial_{t} \gamma_{i}\left(t\right)\right)^{2} dt \to \min_{\Gamma(t), \Theta}$$

$$g(x_t, \theta_i) = \| x_t - \sum_{k=0}^{n} \theta_{ik} \odot \phi_k(t) \|^2$$





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Non-Stationarity in Discrete State Space Time Series



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Historical Circulation Data: weather regimes (Data from the Univ. of East Anglia) 3 atmospherical states considered





Observed Time Series: $\{X_1,\ldots,X_T\}$, $X_t \in s_1,\ldots,s_m$

Markov-Property:

$$P[X_t = s_j | X_1, X_2, \dots, X_{t-1} = s_i] = P[X_t = s_j | X_{t-1} = s_i] = P_{ij}(t)$$

Log-Likelihood:

$$\mathbf{L}(P(t)) = \log P[X_1, \dots, X_T]$$

$$\sum_{j=1}^m P_{ij}(t) = 1, \text{ for all } t, i$$

$$P_{ij}(t) \ge 0, \text{ for all } t, i, j$$

$$= \log P[X_1] + \sum_{i}^m \sum_{j=1}^m \sum_{t \in \{t_{ij}\}} \log P_{ij}(t) \rightarrow \max_{P(t)}$$

Maximization problem is ill-posed => *regularization necessary*





Historical Circulation Data: 27 Lamb regimes (Data from the Univ. of East Anglia) 3 atmospherical states considered

Regularization: global stationarity assumption

<u>ic</u>	anticyclonic	other	cyclonic
anticyclor	0.34	0.40	0.26
other	0.35	0.41	0.24
clonic	0.34	0.41	0.25
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Regularization: local stationarity assumption inside of the window

Gaussian Window:
$$\gamma(t, t_0) = \frac{1}{c} \exp(-\frac{(t-t_0)^2}{\sigma^2})$$

Approximate Log-Likelihood:

$$\mathbf{L}(P(t_0)) \approx \log \mathbf{P}[X_1] + \sum_{i,j=1}^m \sum_{t \in \{t_{ij}\}} \gamma(t,t_0) \log P_{ij}(t_0)$$

Gaussian Kernel Filtering.

$$P_{ij}(t_0) = \frac{\sum_{t \in \{t_{ij}\}} \gamma(t, t_0)}{\sum_{t \in \{t_i\}} \gamma(t, t_0)}$$

Standard statistical methods applicable to calculate the conf. intervals





Regularization: local stationarity assumption inside of the window

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Approximate Log-Likelihood:

$$\mathbf{L}(P(t_0)) \approx \log \mathbf{P}[X_1] + \sum_{i,j=1}^m \sum_{t \in \{t_{ij}\}} \gamma(t,t_0) \log P_{ij}(t_0)$$

Gaussian Kernel Filtering.

$$P_{ij}(t_0) = \frac{\sum_{t \in \{t_{ij}\}} \gamma(t, t_0)}{\sum_{t \in \{t_i\}} \gamma(t, t_0)}$$

Exercise 2: proof that this formula is true



Historical Circulation Data: 27 Lamb regimes (Data from the Univ. of East Anglia) 3 atmospherical states considered

Gaussian Window vs. Homogenous Estimator



other anticyclonic

cyclonic



Historical Circulation Data: 27 Lamb regimes (Data from the Univ. of East Anglia) <u>3 atmospherical states considered</u>

 $\pi(t) P(t) = \pi(t)$







(H. 08, to appear in J. of Atmos. Sci.)

Markovian Trend Model:

$$P(t) = P^{(0)} + P^{(1)}\phi(t), \quad \phi : [1,T] \to (-\infty, +\infty)$$

Log-Likelihood:

$$\sum_{j=1}^{m} \sum_{t \in \{t_{ij}\}} \log \left(P_{ij}^{(0)} + P_{ij}^{(1)} \phi(t) \right) \rightarrow \max_{P^{(0)}, P^{(1)}},$$

$$\sum_{j=1}^{m} P_{ij}^{(0)} = 1,$$

$$\sum_{j=1}^{m} P_{ij}^{(1)} = 0,$$

$$P_{ij}^{(0)} + P_{ij}^{(1)} \sup_{t \in [1,T]} \phi(t) \geq 0, \text{ for all } j,$$

$$P_{ij}^{(0)} + P_{ij}^{(1)} \inf_{t \in [1,T]} \phi(t) \geq 0, \text{ for all } j.$$

Numerics: Nelder-Mead Optimization Algorithm



Historical Circulation Data: 27 Lamb regimes (Data from the Univ. of East Anglia) 3 atmospherical states considered

Polynomial Trend Model





Historical Circulation Data: 27 Lamb regimes (Data from the Univ. of East Anglia) 3 atmospherical states considered

Polynomial Trend Model





Regularized clustering functional:

$$\mathbf{L}^{\epsilon}(\Theta, \Gamma(t), \epsilon^{2}) = \mathbf{L}(\Theta, \Gamma(t)) + \epsilon^{2} \sum_{i=1}^{\mathbf{K}} \int_{0}^{T} \left(\partial_{t} \gamma_{i}\left(t\right)\right)^{2} dt \to \min_{\Gamma(t), \Theta}$$

$$\mathbf{L}(\Theta, \Gamma(t)) = \int_0^T \sum_{i=1}^{\mathbf{K}} \gamma_i(t) g(x_t, \theta_i) \to \min_{\Gamma(t), \Theta},$$





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$$\mathbf{L}^{\epsilon}(\Theta, \Gamma(t), \epsilon^{2}) = \mathbf{L}(\Theta, \Gamma(t)) + \epsilon^{2} \sum_{i=1}^{\mathbf{K}} \int_{0}^{T} \left(\partial_{t} \gamma_{i}\left(t\right)\right)^{2} dt \to \min_{\Gamma(t), \Theta}$$

$$\mathbf{L}(\Theta, \Gamma(t)) = \int_0^T \sum_{i=1}^K \gamma_i(t) g\left(x_t, \theta_i\right) \to \min_{\Gamma(t), \Theta}$$
$$x_t : X_t \to X_{t+1}$$
$$g(x_t, P^i) = -\log P^i_{X_t X_{t+1}}$$

$$\sum_{i=1}^{\mathbf{K}} \gamma_i(t) = 1, \quad \forall t \in [0, T]$$
$$\gamma_i(t) \geq 0, \quad \forall t \in [0, T], i = 1, \dots, \mathbf{K}$$



(H. 08, to appear in J. of Atmos. Sci.) Regularized clustering functional: $\mathbf{L}^{\epsilon}(\Theta, \Gamma(t), \epsilon^{2}) = \mathbf{L}(\Theta, \Gamma(t)) + \epsilon^{2} \sum_{i=1}^{\infty} \int_{0}^{1} \left(\partial_{t} \gamma_{i}\left(t\right)\right)^{2} dt \to \min_{\Gamma(t), \Theta}$ $x_t: X_t \to X_{t+1}$ $g(x_t, P^i) = -\log P^i_{X_t X_{t+1}}$ rative Subspace $\tilde{\mathbf{L}}^{\epsilon} = \sum_{i=1} \left[a^{\mathbf{T}}(\theta_i) \bar{\gamma}_i + \epsilon^2 \bar{\gamma}_i^{\mathbf{T}} \mathbf{H} \bar{\gamma}_i \right] \to \min_{\bar{\gamma}_i, \Theta}$ QP can be used
$$\begin{split} \sum_{i=1}^{\mathbf{K}} \tilde{\gamma}_i^{(k+1)} &= 1, \quad \forall k = 1, \dots, N, \\ \tilde{\gamma}_i^{(k+1)} &\geq 0, \quad \forall k = 1, \dots, N; i = 1, \dots, \mathbf{K}. \end{split}$$
Minimizatic



Historical Circulation Data: 28 Lamb regimes (Data from the Univ. of East Anglia) 3 atmospherical states considered



 $\mathbf{L}\left(\gamma_{i}(t)\right)$

mit

 X_{t+1}

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Historical Circulation Data: 28 Lamb regimes (Data from the Univ. of East Anglia) 3 atmospherical states considered

Circulation Patterns for UK(1945-2007)

Gaussian Kernel Estimator

FEM-Clustering vs. Single Trend



Example: analysis of hystorical temperatures (1947-2007)



Temperature Data Analysis



Global Historical Temperatures

Historical Temperatures in Europe

Deviation from the Mean Temperature





Data from Boulder Center, USA

Data from Ch. Franzke, BAS Cambridge

Temperature Data on the 2D-Grid (1947-2007)



Temperature Data in Europe: 29x20 grid

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Seasonal Cycle Eliminated



$$\mathbf{L}(\gamma_{i}(t), \mathbf{T}_{i}, \mu_{i}) \rightarrow \min : \mathsf{K=1}$$

$$g(x_{t}, \theta_{i}) = || x_{t} - \sum_{k=0}^{\mathcal{R}} \theta_{ik} \odot \phi_{k}(t) ||^{2}$$
Linear Regression

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Temperature Data in Europe: 29x20 grid

Seasonal Cycle Eliminated







Temperature Data in Europe: 29x20 grid

∆ T(01-Jan-1947)

Cluster 1



<u>Cluster 2</u>



Cluster 3







Temperature Data in Europe: 29x20 grid

∆ T(31-Dec-2007)

Cluster 1



Cluster 2



Cluster 3







Temperature Data in Europe: 29x20 grid

∆ T(01-Jan-1947)

Cluster 1



<u>Cluster 2</u>



Cluster 3







Temperature Data in Europe: 29x20 grid

∆ T(31-Dec-2007)

Cluster 1



<u>Cluster 2</u>



Cluster 3







Temperature Data in Europe: 29x20 grid







Global Historical Temperature Data (80x120 grid, daily values 1947-2007) (Data from the NCAR, Boulder, US)

∆T(Start)







NCEPNCAR/air.1948.2007.1000.mat







Global Historical Temperature Data (80x120 grid, daily values 1947-2007) (Data from the NCAR, Boulder, US)



Take-Home-Messages

- 1. Different methods for analysis of non-stationary data were presented
- 2. One can use similar approaches in both continuous and discrete data analysis
- 3. Issue of non-stationarity is important in analysis of historical time series







Thank you for attention!